

3D Computer Vision

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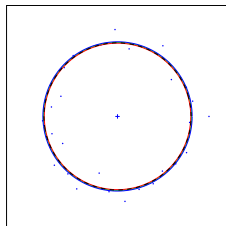
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Open Informatics Master's Course

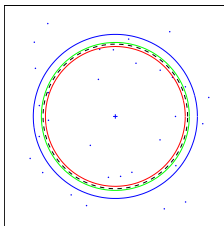
Circle Fitting: Some Results

medium radial noise



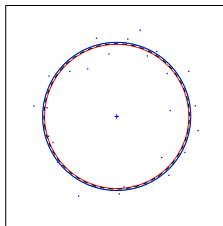
opt: 1.8, Smp: 1.9, dir: 2.3

big radial noise



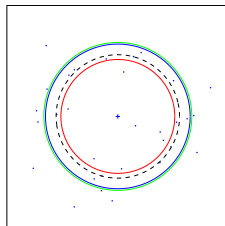
1.6, 1.8, 2.6

medium isotropic noise



1.8, 2.0, 2.2

big isotropic noise



1.6, 2.0, 2.4

mean ranks over 10000 random trials with $k = 32$ samples

- green – ground truth
- red – Sampson error e minimizer
- blue – direct radial error ϵ minimizer
- black – optimal estimator for isotropic error

optimal estimator for isotropic error (black, dashed):

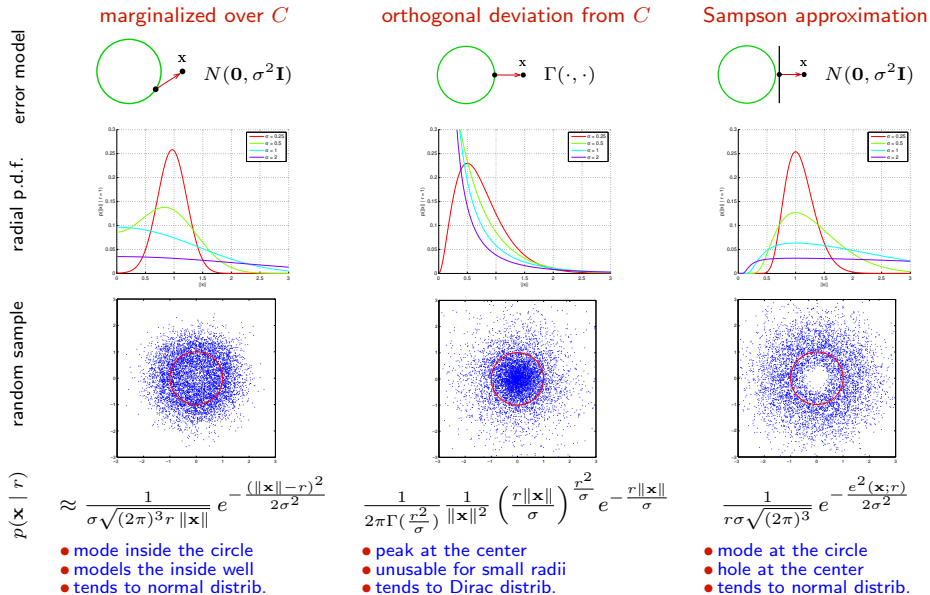
$$r \approx \frac{3}{4k} \sum_{i=1}^k \|\mathbf{x}_i\| + \sqrt{\left(\frac{3}{4k} \sum_{i=1}^k \|\mathbf{x}_i\| \right)^2 - \frac{1}{2k} \sum_{i=1}^k \|\mathbf{x}_i\|^2}$$

which method is better?

- error should model noise, radial noise and isotropic noise behave differently
- ground truth: Normally distributed isotropic error, Gamma-distributed radial error
- Sampson: better for the radial distribution model; Direct: better for the isotropic model
- no matter how corrected, the algebraic error minimizer is not an unbiased parameter estimator
Cramér-Rao bound tells us how close one can get with unbiased estimator and given k

Discussion: On The Art of Probabilistic Model Design...

- a few models for fitting zero-centered circle C of radius r to points in \mathbb{R}^2



► Sampson Error for Fundamental Matrix Manifold

The epipolar algebraic error is

$$\varepsilon_i(\mathbf{F}) = \underline{\mathbf{y}}_i^\top \mathbf{F} \underline{\mathbf{x}}_i, \quad \mathbf{x}_i = (u_i^1, v_i^1), \quad \mathbf{y}_i = (u_i^2, v_i^2), \quad \varepsilon_i \in \mathbb{R}$$

Let $\mathbf{F} = [\mathbf{F}_1 \quad \mathbf{F}_2 \quad \mathbf{F}_3]$ (per columns) = $\begin{bmatrix} (\mathbf{F}^1)^\top \\ (\mathbf{F}^2)^\top \\ (\mathbf{F}^3)^\top \end{bmatrix}$ (per rows), $\mathbf{S} = \begin{bmatrix} 1 & 0 & 0 \\ 0 & 1 & 0 \end{bmatrix}$, then

Sampson

$$\mathbf{J}_i(\mathbf{F}) = \left[\frac{\partial \varepsilon_i(\mathbf{F})}{\partial u_i^1}, \frac{\partial \varepsilon_i(\mathbf{F})}{\partial v_i^1}, \frac{\partial \varepsilon_i(\mathbf{F})}{\partial u_i^2}, \frac{\partial \varepsilon_i(\mathbf{F})}{\partial v_i^2} \right] \quad \mathbf{J}_i \in \mathbb{R}^{1,4} \quad \text{derivatives over point coordinates}$$

$$= \left[(\mathbf{F}_1)^\top \underline{\mathbf{y}}_i, (\mathbf{F}_2)^\top \underline{\mathbf{y}}_i, (\mathbf{F}^1)^\top \underline{\mathbf{x}}_i, (\mathbf{F}^2)^\top \underline{\mathbf{x}}_i \right] = \begin{bmatrix} \mathbf{S}\mathbf{F}^\top \underline{\mathbf{y}}_i \\ \mathbf{S}\mathbf{F} \underline{\mathbf{x}}_i \end{bmatrix}^\top$$

$$\mathbf{e}_i(\mathbf{F}) = -\frac{\mathbf{J}_i(\mathbf{F}) \varepsilon_i(\mathbf{F})}{\|\mathbf{J}_i(\mathbf{F})\|^2} \quad \mathbf{e}_i(\mathbf{F}) \in \mathbb{R}^4 \quad \text{Sampson error vector}$$

$$e_i(\mathbf{F}) \stackrel{\text{def}}{=} \|\mathbf{e}_i(\mathbf{F})\| = \frac{\varepsilon_i(\mathbf{F})}{\|\mathbf{J}_i(\mathbf{F})\|} = \frac{\underline{\mathbf{y}}_i^\top \mathbf{F} \underline{\mathbf{x}}_i}{\sqrt{\|\mathbf{S}\mathbf{F} \underline{\mathbf{x}}_i\|^2 + \|\mathbf{S}\mathbf{F}^\top \underline{\mathbf{y}}_i\|^2}} \quad e_i(\mathbf{F}) \in \mathbb{R} \quad \text{scalar Sampson error}$$

- Sampson error 'normalizes' the algebraic error
- automatically copes with multiplicative factors $\mathbf{F} \mapsto \lambda \mathbf{F}$
- actual optimization not yet covered →110

► Back to Triangulation: The Golden Standard Method

Given \mathbf{P}_1 , \mathbf{P}_2 and a correspondence $x \leftrightarrow y$, look for 3D point \mathbf{X} projecting to x and y →89

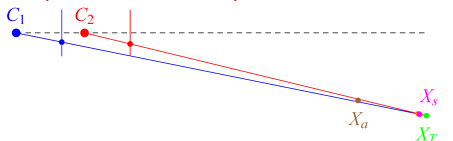
Idea:

1. if not given, compute \mathbf{F} from \mathbf{P}_1 , \mathbf{P}_2 , e.g. $\mathbf{F} = (\mathbf{Q}_1 \mathbf{Q}_2^{-1})^\top [\mathbf{q}_1 - (\mathbf{Q}_1 \mathbf{Q}_2^{-1}) \mathbf{q}_2]_{\times}$
2. correct the measurement by the linear estimate of the correction vector →101

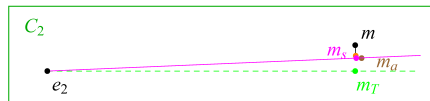
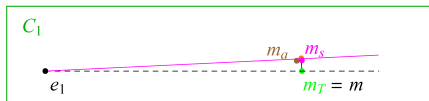
$$\begin{bmatrix} \hat{u}^1 \\ \hat{v}^1 \\ \hat{u}^2 \\ \hat{v}^2 \end{bmatrix} \approx \begin{bmatrix} u^1 \\ v^1 \\ u^2 \\ v^2 \end{bmatrix} - \frac{\varepsilon}{\|\mathbf{J}\|^2} \mathbf{J}^\top = \begin{bmatrix} u^1 \\ v^1 \\ u^2 \\ v^2 \end{bmatrix} - \frac{\underline{\mathbf{y}}^\top \mathbf{F} \underline{\mathbf{x}}}{\|\mathbf{S} \mathbf{F} \underline{\mathbf{x}}\|^2 + \|\mathbf{S} \mathbf{F}^\top \underline{\mathbf{y}}\|^2} \begin{bmatrix} (\mathbf{F}_1)^\top \underline{\mathbf{y}} \\ (\mathbf{F}_2)^\top \underline{\mathbf{y}} \\ (\mathbf{F}^1)^\top \underline{\mathbf{x}} \\ (\mathbf{F}^2)^\top \underline{\mathbf{x}} \end{bmatrix}$$

3. use the SVD triangulation algorithm with numerical conditioning →90

Ex (cont'd from →93):



- X_T – noiseless ground truth position
- – reprojection error minimizer
- X_s – Sampson-corrected algebraic error minimizer
- X_a – algebraic error minimizer
- m – measurement (m_T with noise in v^2)



► Back to Fundamental Matrix Estimation

Goal: Given a set $X = \{(x_i, y_i)\}_{i=1}^k$ of $k \gg 7$ inlier correspondences, compute a statistically efficient estimate for fundamental matrix \mathbf{F} .

What we have so far

- 7-point algorithm for \mathbf{F} (5-point algorithm for \mathbf{E}) → 84
- definition of Sampson error per correspondence $e_i(\mathbf{F} \mid x_i, y_i)$ → 105
- triangulation requiring an optimal \mathbf{F}

What we need

- correspondence recognition
- an optimization algorithm for many ($k \gg 7$) correspondences

$$\mathbf{F}^* = \arg \min_{\mathbf{F}} \sum_{i=1}^k e_i^2(\mathbf{F} \mid X)$$

- the 7-point estimate is a good starting point \mathbf{F}_0

Levenberg-Marquardt (LM) Iterative Optimization in a Nutshell

Consider error function $\mathbf{e}_i(\boldsymbol{\theta}) = f(\mathbf{x}_i, \mathbf{y}_i, \boldsymbol{\theta}) \in \mathbb{R}^m$, with $\mathbf{x}_i, \mathbf{y}_i$ given, $\boldsymbol{\theta} \in \mathbb{R}^q$ unknown
 $\boldsymbol{\theta} = \mathbf{F}$, $q = 9$, $m = 1$ for f.m. estimation

Our goal: $\boldsymbol{\theta}^* = \arg \min_{\boldsymbol{\theta}} \sum_{i=1}^k \|\mathbf{e}_i(\boldsymbol{\theta})\|^2$

Idea 1 (Gauss-Newton approximation): proceed iteratively for $s = 0, 1, 2, \dots$

$$\boldsymbol{\theta}^{s+1} := \boldsymbol{\theta}^s + \mathbf{d}_s, \quad \text{where } \mathbf{d}_s = \arg \min_{\mathbf{d}} \sum_{i=1}^k \|\mathbf{e}_i(\boldsymbol{\theta}^s + \mathbf{d})\|^2 \quad (19)$$

$$\mathbf{e}_i(\boldsymbol{\theta}^s + \mathbf{d}) \approx \mathbf{e}_i(\boldsymbol{\theta}^s) + \mathbf{L}_i \mathbf{d},$$

$$(\mathbf{L}_i)_{jl} = \frac{\partial (\mathbf{e}_i(\boldsymbol{\theta}))_j}{\partial (\boldsymbol{\theta})_l}, \quad \mathbf{L}_i \in \mathbb{R}^{m,q} \quad \text{typically a long matrix, } m \ll q$$

Then the solution to Problem (19) is a set of 'normal' eqs

$$-\underbrace{\sum_{i=1}^k \mathbf{L}_i^\top \mathbf{e}_i(\boldsymbol{\theta}^s)}_{\mathbf{e} \in \mathbb{R}^{q,1}} = \underbrace{\left(\sum_{i=1}^k \mathbf{L}_i^\top \mathbf{L}_i \right)}_{\mathbf{L} \in \mathbb{R}^{q,q}} \mathbf{d}_s, \quad (20)$$

- \mathbf{d}_s can be solved for by Gaussian elimination using Choleski decomposition of \mathbf{L}
 \mathbf{L} symmetric PSD \Rightarrow use Choleski, almost $2\times$ faster than Gauss-Seidel, see bundle adjustment $\rightarrow 141$
- beware of rank deficiency in \mathbf{L} when k is small
- such updates do not lead to stable convergence \rightarrow ideas of Levenberg and Marquardt

Idea 2 (Levenberg): replace $\sum_i \mathbf{L}_i^\top \mathbf{L}_i$ with $\sum_i \mathbf{L}_i^\top \mathbf{L}_i + \lambda \mathbf{I}$ for some damping factor $\lambda \geq 0$

Idea 3 (Marquardt): replace $\lambda \mathbf{I}$ with $\lambda \sum_i \text{diag}(\mathbf{L}_i^\top \mathbf{L}_i)$ to adapt to local curvature:

$$-\sum_{i=1}^k \mathbf{L}_i^\top \mathbf{e}_i(\boldsymbol{\theta}^s) = \left(\sum_{i=1}^k (\mathbf{L}_i^\top \mathbf{L}_i + \lambda \text{diag}(\mathbf{L}_i^\top \mathbf{L}_i)) \right) \mathbf{d}_s$$

Idea 4 (Marquardt): adaptive λ small $\lambda \rightarrow$ Gauss-Newton, large $\lambda \rightarrow$ gradient descend

1. choose $\lambda \approx 10^{-3}$ and compute \mathbf{d}_s
2. if $\sum_i \|\mathbf{e}_i(\boldsymbol{\theta}^s + \mathbf{d}_s)\|^2 < \sum_i \|\mathbf{e}_i(\boldsymbol{\theta}^s)\|^2$ then accept \mathbf{d}_s and set $\lambda := \lambda/10$, $s := s + 1$
3. otherwise set $\lambda := 10\lambda$ and recompute \mathbf{d}_s

- sometimes different constants are needed for the 10 and 10^{-3}
- note that $\mathbf{L}_i \in \mathbb{R}^{m,q}$ (long matrix) but each contribution $\mathbf{L}_i^\top \mathbf{L}_i$ is a square singular $q \times q$ matrix (always singular for $k < q$)
- λ helps avoid the consequences of gauge freedom $\rightarrow 143$
- error can be made robust to outliers $\rightarrow 113$
- we have approximated the least squares Hessian by ignoring second derivatives of the error function (Gauss-Newton approximation) See [Triggs et al. 1999, Sec. 4.3]
- modern variants of LM are Trust Region methods

LM with Sampson Error for Fundamental Matrix Estimation

Sampson (derived by linearization over point coordinates u^1, v^1, u^2, v^2)

$$e_i(\mathbf{F}) = \frac{\varepsilon_i}{\|\mathbf{J}_i\|} = \frac{\underline{\mathbf{y}}_i^\top \mathbf{F} \underline{\mathbf{x}}_i}{\sqrt{\|\mathbf{S} \mathbf{F} \underline{\mathbf{x}}_i\|^2 + \|\mathbf{S} \mathbf{F}^\top \underline{\mathbf{y}}_i\|^2}} \quad \text{where} \quad \mathbf{S} = \begin{bmatrix} 1 & 0 & 0 \\ 0 & 1 & 0 \end{bmatrix}$$

LM (by linearization over parameters \mathbf{F})

$$\mathbf{L}_i = \frac{\partial e_i(\mathbf{F})}{\partial \mathbf{F}} = \dots = \frac{1}{2\|\mathbf{J}_i\|} \left[\left(\underline{\mathbf{y}}_i - \frac{2e_i}{\|\mathbf{J}_i\|} \mathbf{S} \mathbf{F} \underline{\mathbf{x}}_i \right) \underline{\mathbf{x}}_i^\top + \underline{\mathbf{y}}_i \left(\underline{\mathbf{x}}_i - \frac{2e_i}{\|\mathbf{J}_i\|} \mathbf{S} \mathbf{F}^\top \underline{\mathbf{y}}_i \right)^\top \right] \quad (21)$$

- \mathbf{L}_i in (21) is a 3×3 matrix, must be reshaped to dimension-9 vector $\text{vec}(\mathbf{L}_i)$ to be used in LM
- $\underline{\mathbf{x}}_i$ and $\underline{\mathbf{y}}_i$ in Sampson error are normalized to unit homogeneous coordinate (21) relies on this
- reinforce rank $\mathbf{F} = 2$ after each LM update to stay on the fundamental matrix manifold and $\|\mathbf{F}\| = 1$ to avoid gauge freedom by SVD $\rightarrow 111$
- LM linearization could be done by numerical differentiation (we have a small dimension here)

► Local Optimization for Fundamental Matrix Estimation

Summary so far

- Given a set $X = \{(x_i, y_i)\}_{i=1}^k$ of $k \gg 7$ inlier correspondences, compute a statistically efficient estimate for fundamental matrix \mathbf{F} .
 - Find the conditioned ($\rightarrow 92$) 7-point \mathbf{F}_0 ($\rightarrow 84$) from a suitable 7-tuple
 - Improve the \mathbf{F}_0^* using the LM optimization ($\rightarrow 108-109$) and the Sampson error ($\rightarrow 110$) on all inliers, reinforce rank-2, unit-norm \mathbf{F}_k^* after each LM iteration using SVD

Partial conceptualization

- inlier = correspondence
- outlier = non-correspondence
- binary inlier/outlier labels are hidden
- we can get their likely estimate only, with respect to a model

We are not yet done

- if there are no wrong correspondences (mismatches, outliers), this gives a local optimum given the 7-point initial estimate
- the algorithm breaks under contamination of (inlier) correspondences by outliers
- the full problem involves finding the inliers!
- in addition, we need a mechanism for jumping out of local minima (and exploring the space of all fundamental matrices)

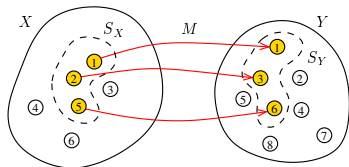
► The Full Problem of Matching and Fundamental Matrix Estimation

Problem: Given image keypoint sets $X = \{x_i\}_{i=1}^m$ and $Y = \{y_j\}_{j=1}^n$ and their descriptors D , find the most probable

1. inlier keypoints $S_X \subseteq X$, $S_Y \subseteq Y$
2. one-to-one perfect matching $M: S_X \rightarrow S_Y$
3. fundamental matrix \mathbf{F} such that $\text{rank } \mathbf{F} = 2$
4. such that for each $x_i \in S_X$ and $y_j = M(x_i)$ it is probable that
 - a) the image descriptor $D(x_i)$ is similar to $D(y_j)$, and
 - b) the total reprojection error $E = \sum_{ij} e_{ij}^2(\mathbf{F})$ is small
5. inlier-outlier and outlier-outlier matches are improbable

perfect matching: 1-factor of the bipartite graph

note a slight change in notation: e_{ij}



$M:$

| | Y | | | | | | | |
|---|---|---|---|---|---|---|---|---|
| X | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 |
| 1 | 1 | | | | | | | |
| 2 | | | 1 | | | | | |
| 3 | | | | | | | | |
| 4 | | | | | | | | |
| 5 | | | | | | 1 | | |
| 6 | | | | | | | | |

□ = 0
 ■ = 1 (matched)

$$(M^*, \mathbf{F}^*) = \arg \max_{M, \mathbf{F}} p(E, D, \mathbf{F} \mid M) P(M) \quad (22)$$

- probabilistic model: an efficient language for problem formulation it also unifies 4.a and 4.b
- the (22) is a Bayesian probabilistic model there is a constant number of random variables!
- binary matching table $M_{ij} \in \{0, 1\}$ of fixed size $m \times n$
 - each row/column contains at most one unity
 - zero rows/columns correspond to unmatched point x_i/y_j

Deriving A Robust Matching Model by Approximate Marginalization

For algorithmic efficiency, instead of $(M^*, \mathbf{F}^*) = \arg \max_{M, \mathbf{F}} p(E, D, \mathbf{F} | M) P(M)$ solve

$$\mathbf{F}^* = \arg \max_{\mathbf{F}} p(E, D, \mathbf{F}) \quad (23)$$

by marginalization of $p(E, D, \mathbf{F} | M) P(M)$ over \mathcal{M} s.t. $M \in \mathcal{M}$ this changes the problem!
drop the assumption that M is a 1:1 matching, assume correspondence-wise independence:

$$p(E, D, \mathbf{F} | M) P(M) = \prod_{i=1}^m \prod_{j=1}^n p_e(e_{ij}, d_{ij}, \mathbf{F} | m_{ij}) P(m_{ij})$$

- e_{ij} represents (reprojection) error for match $x_i \leftrightarrow y_j$: e.g. $e_{ij}(x_i, y_j, \mathbf{F})$
- d_{ij} represents descriptor similarity for match $x_i \leftrightarrow y_j$: e.g. $d_{ij} = \|\mathbf{d}(x_i) - \mathbf{d}(y_j)\|$

Approximate marginalization:

take all the 2^{mn} terms in place of M

$$\begin{aligned} p(E, D, \mathbf{F}) &\approx \sum_{m_{11} \in \{0,1\}} \sum_{m_{12}} \cdots \sum_{m_{mn}} p(E, D, \mathbf{F} | M) P(M) = \\ &= \sum_{m_{11}} \cdots \sum_{m_{mn}} \prod_{i=1}^m \prod_{j=1}^n p_e(e_{ij}, d_{ij}, \mathbf{F} | m_{ij}) P(m_{ij}) = \overset{*}{\dots} \overset{!}{=} \\ &= \prod_{i=1}^m \prod_{j=1}^n \underbrace{\sum_{m_{ij} \in \{0,1\}} p_e(e_{ij}, d_{ij}, \mathbf{F} | m_{ij}) P(m_{ij})}_{\text{we will continue with this term}} \end{aligned}$$

Thank You

